

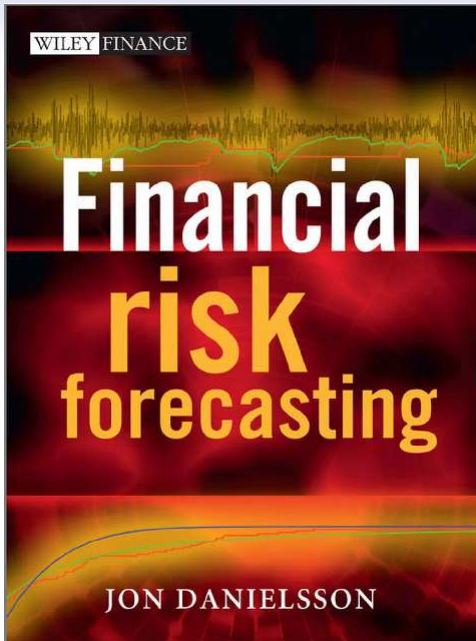
Financial Risk Forecasting

Chapter 9

Extreme Value Theory

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London School of Economics

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The Focus of This Chapter

- Basic introduction to extreme value theory (EVT)
- Asset returns and fat tails
- Applying EVT
- Aggregation and convolution
- Time dependence

Notation

ι	Tail index
$\xi = 1/\iota$	Shape parameter
M_T	Maximum of X
C_T	Number of observations in the tail
u	Threshold value
ψ	Extremal index

Extreme Value Theory

Types of Tails

- In this book, we follow the convention of EVT being presented in terms of the *upper tails* (ie *positive observations*)
- In most risk analysis we are concerned with the *negative observations* in the lower tails, hence to follow the convention, we can *pre-multiply returns by -1*
- Note, the upper and lower tails do not need to have the same thickness or shape

Extreme Value Distributions

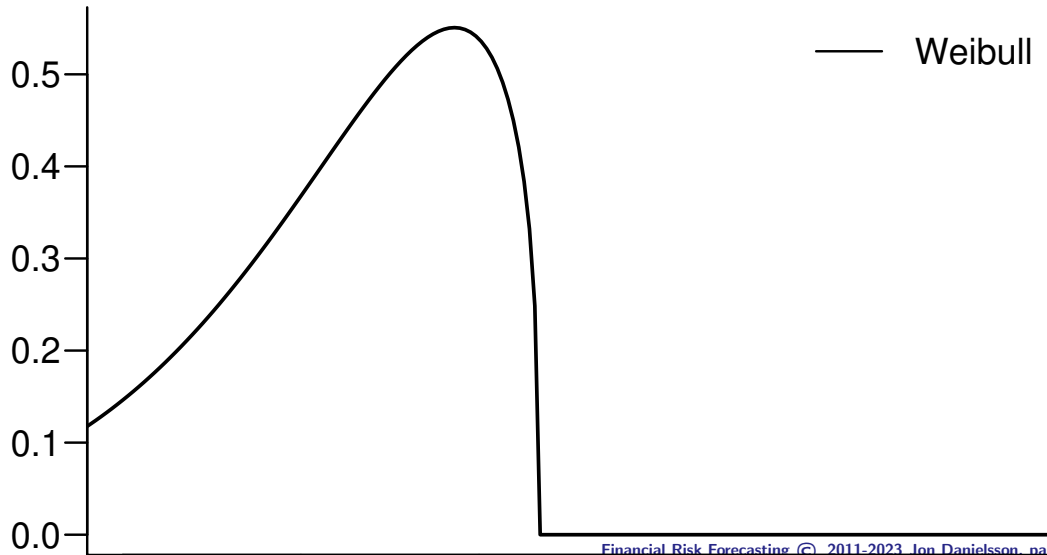
- In most risk applications, we do not need to focus on the entire distribution
- The main result of EVT states that the tails of all distributions fall into one of three categories, regardless of the overall shape of the distribution
 - See next slide for the three distributions
- Note, this is true given the distribution of an asset return does not change over time

Weibull Thin tails where the distribution has a finite endpoint (eg the distribution of mortality and insurance/re-insurance claims)

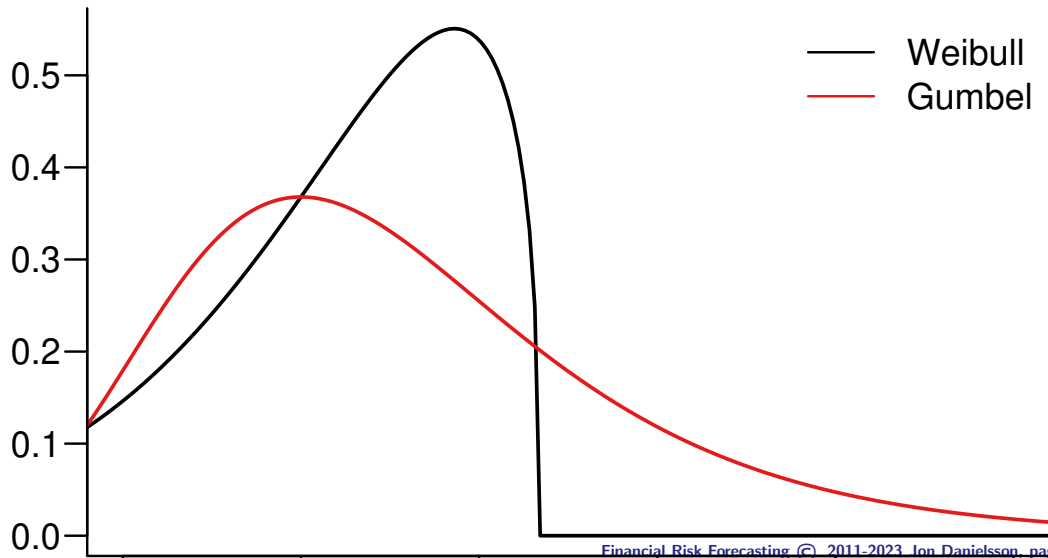
Gumbel Tails decline exponentially (eg the normal and log-normal distributions)

Fréchet Tails decline by a *power law*; such tails are known as “fat tails” (eg the Student-t and Pareto distributions)

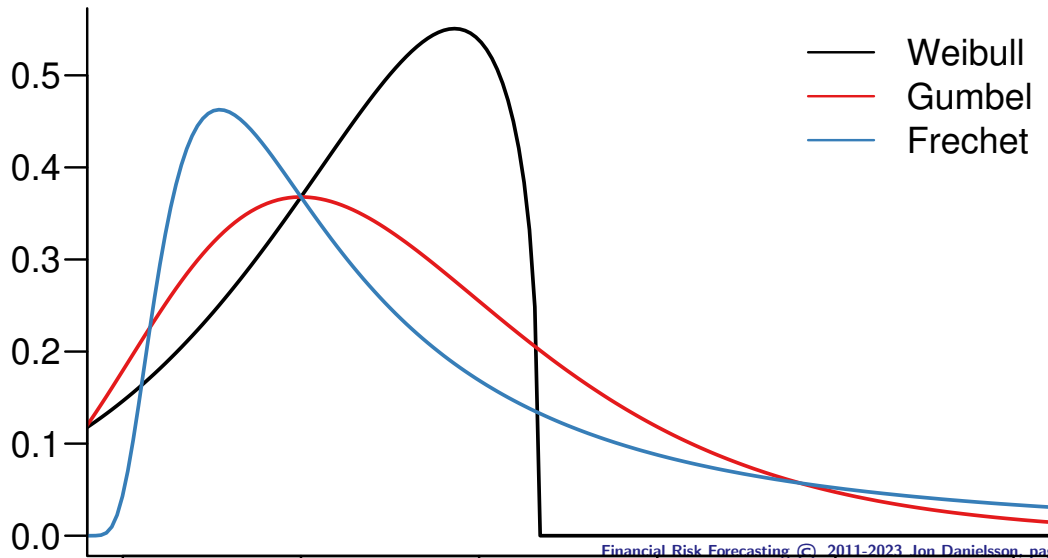
Extreme Value Distributions



Extreme Value Distributions



Extreme Value Distributions



Fréchet distribution

- From the last slide, the Weibull clearly has a finite endpoint
- And the Fréchet tail is thicker than the Gumbel's
- In most applications in finance, we know that returns are fat tailed
- Hence we limit our attention to the Fréchet case

Generalised Extreme Value Distribution

- The Fisher and Tippett (1928) and Gnedenko (1943) theorems are the fundamental results in EVT
- The theorems state that the maximum of a sample of properly normalised IID random variables *converges in distribution* to one of the three possible distributions: the Weibull, Gumbel or the Fréchet

Generalised Extreme Value Distribution

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- The theorems state that the maximum of a sample of properly normalised IID random variables *converges in distribution* to one of the three possible distributions: the Weibull, Gumbel or the Fréchet
- An alternative way of stating this is in terms of the maximum domain of attraction(MDA)
- MDA is the set of limiting distributions for the properly normalised maxima as the sample size goes to infinity

Fisher-Tippett and Gnedenko Theorems

- Let X_1, X_2, \dots, X_T denote IID random variables (RVs) and the term M_T indicate maxima in sample of size T
- The *standardised distribution* of maxima, M_T , is

$$\lim_{T \rightarrow \infty} \Pr \left\{ \frac{M_T - a_T}{b_T} \leq x \right\} = H(x)$$

where the constants a_T and $b_T > 0$ exist and are defined as $a_T = T\mathbb{E}(X_1)$ and $b_T = \sqrt{\text{Var}(X_1)}$

Fisher-Tippett and Gnedenko Theorems

- Then the limiting distribution, $H(\cdot)$, of the maxima as the *generalised extreme value (GEV)* distribution is

$$H_{\xi}(x) = \begin{cases} \exp \left\{ -(1 + \xi x)^{-\frac{1}{\xi}} \right\}, & \xi \neq 0 \\ \exp \{ -\exp(-x) \}, & \xi = 0 \end{cases}$$

Limiting Distribution $H_\xi(\cdot)$

- Depending on the value of ξ , $H_\xi(\cdot)$ becomes one of the three distributions:
 - if $\xi > 0$, $H_\xi(\cdot)$ is the **Fréchet**
 - if $\xi < 0$, $H_\xi(\cdot)$ is the **Weibull**
 - if $\xi = 0$, $H_\xi(\cdot)$ is the **Gumbel**



Asset Returns and Fat Tails

Fat Tails

- The term *“fat tails”* can have several meanings, the most common being *“extreme outcomes occur more frequently than predicted by normal distribution”*
- While such a statement might make intuitive sense, it has little mathematical rigor as stated
- The most frequent definition one may encounter is Kurtosis, but it is not always accurate at indicating the presence of fat tails ($\kappa > 3$)
- This is because kurtosis is more concerned with the sides of the distribution rather than the *heaviness of tails*

A Formal Definition of Fat Tails

- The formal definition of fat tails comes from *regular variation*

Regular variation A random variable, X , with distribution $F(\cdot)$ has fat tails if it varies regularly at infinity; that is there exists a positive constant ι such that:

$$\lim_{t \rightarrow \infty} \frac{1 - F(tx)}{1 - F(t)} = x^{-\iota}, \quad \forall x > 0, \iota > 0$$

Tail Distributions

- In the fat-tailed case, the tail distribution is Fréchet:

$$H(x) = \exp(-x^{-\iota})$$

Lemma *A random variable X has regular variation at infinity (ie has fat tails) if and only if its distribution function F satisfies the following condition:*

$$1 - F(x) = \mathbb{P}\{X > x\} = Ax^{-\iota} + o(x^{-\iota})$$

for positive constant A , when $x \rightarrow \infty$

Tail Distributions

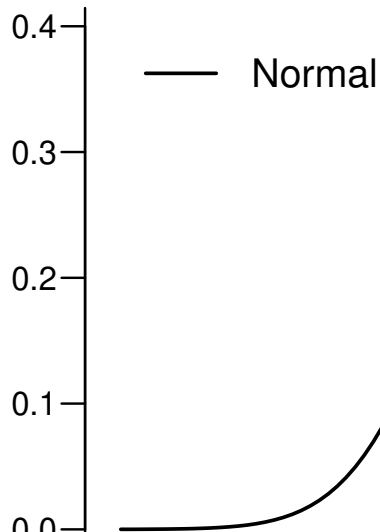
- The expression $o(x^{-\iota})$ is the *remainder term* of the Taylor-expansion of $\Pr\{X > x\}$, it consists of terms of the type Cx^{-j} for constant C and $j > \iota$
- As $x \rightarrow \infty$, the tails are asymptotically Pareto- distributed:

$$F(x) \approx 1 - Ax^{-\iota}$$

where $A > 0$; $\iota > 0$; and $\forall x > A^{1/\iota}$

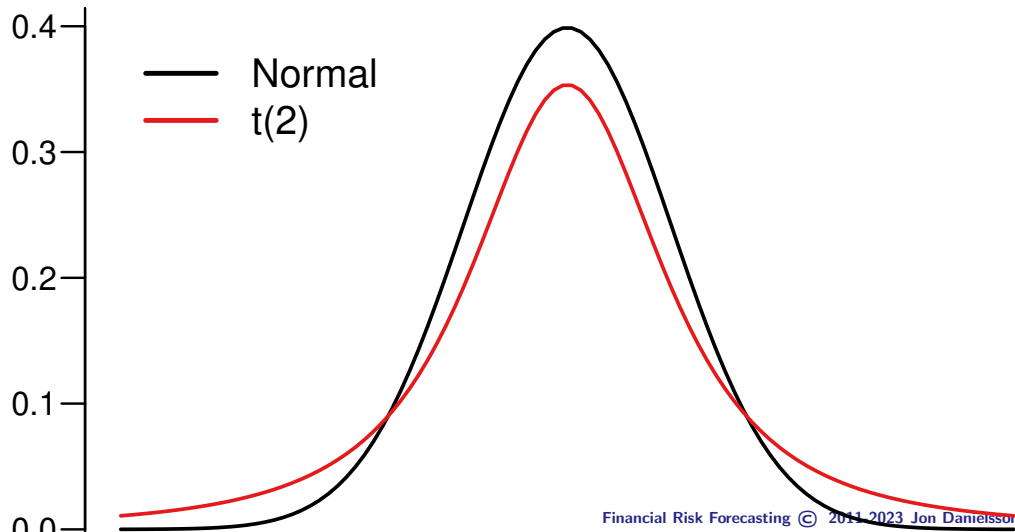
Normal and Fat Distributions

Normal and Student-t densities



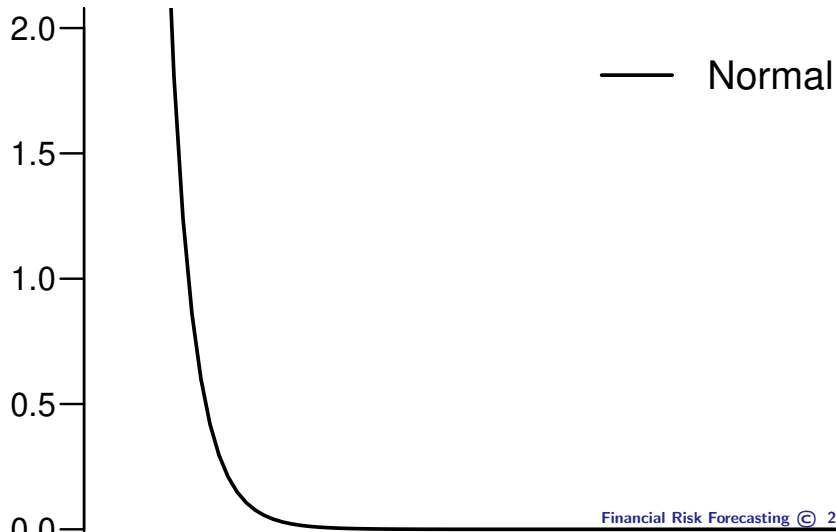
Normal and Fat Distributions

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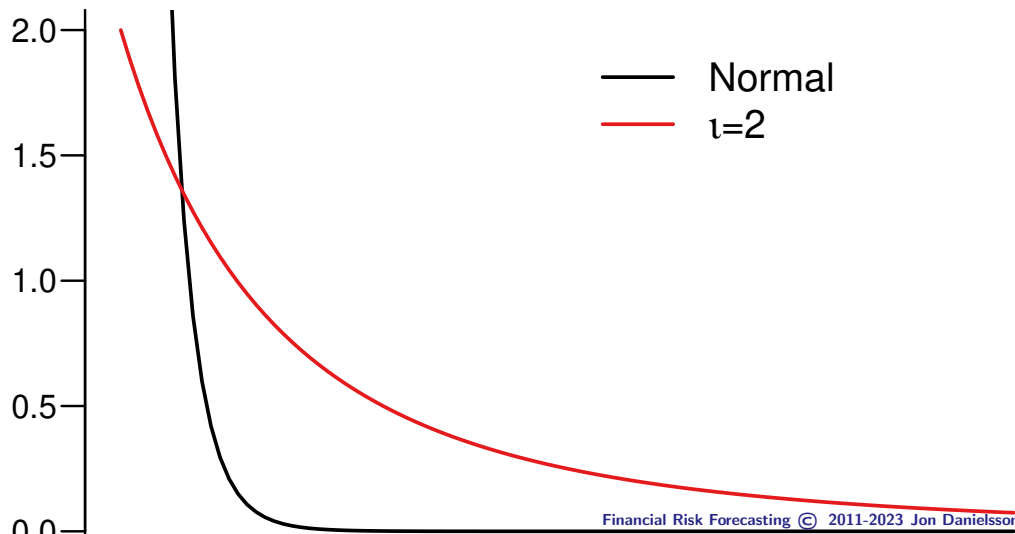
Normal and Fat Distributions

Pareto tails



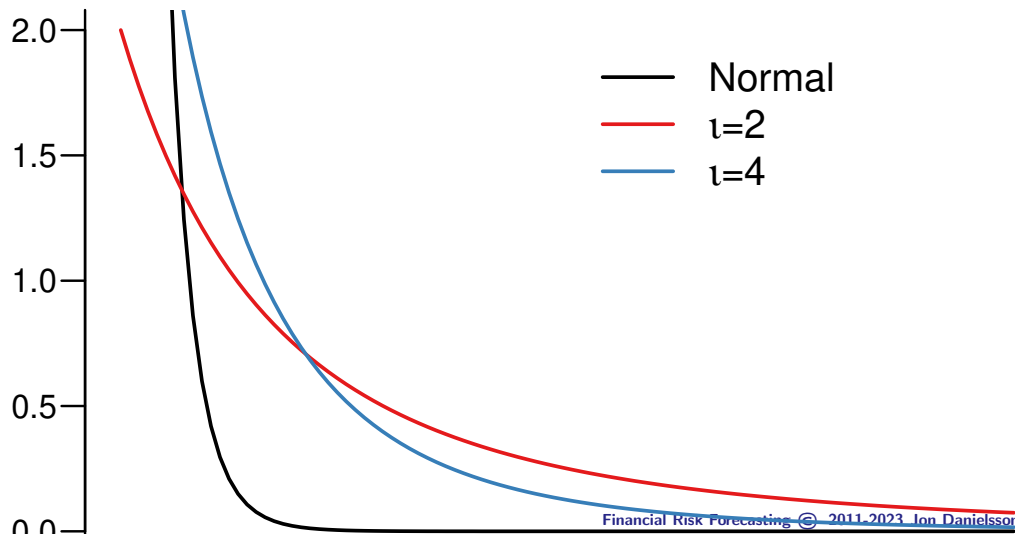
Normal and Fat Distributions

Pareto tails



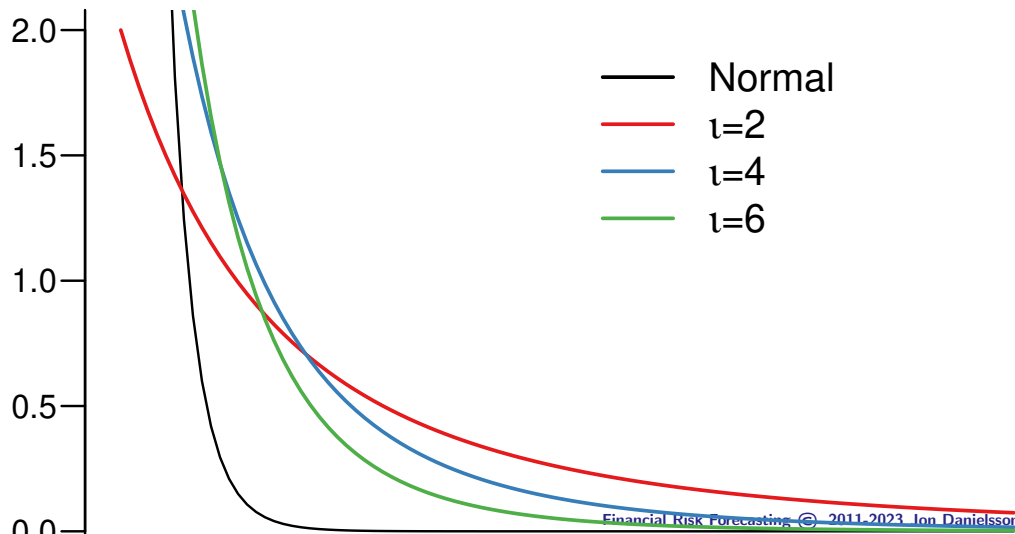
Normal and Fat Distributions

Pareto tails



Normal and Fat Distributions

Pareto tails



Normal and Fat Distributions

- The definition demonstrates that fat tails are defined by how rapidly the tails of the distribution decline as we approach infinity
- As the tails become thicker, we detect increasingly large observations that impact the calculation of moments:

$$E(X^m) = \int x^m f(x) dx$$

- If $E(X^m)$ exists for all positive m , such as for the normal distribution, the definition of *regular variation* implies that moments $m \geq \iota$ are not defined for fat-tailed data

Applying EVT

Implementing EVT in Practice

Two main approaches:

1. Block maxima
2. Peaks over thresholds (POT)

Block Maxima Approach

- This approach follows directly from the regular variation definition where we estimate the GEV by dividing the sample into blocks and using the maxima in each block for estimation
- The procedure is rather wasteful of data and a relatively large sample is needed for accurate estimate

Peaks Over Thresholds Approach

- This approach is generally preferred and forms the basis of our approach below
- It is based on models for all large observations that exceed a high threshold and hence makes better use of data on extreme values
- There are two common approaches to POT:
 1. Fully parametric models (eg *the Generalised Pareto distribution or GPD*)
 2. Semi-parametric models (eg *the Hill estimator*)

Generalised Pareto Distribution

- Consider a random variable X , fix a threshold u and focus on the *positive part of $X - u$*
- The distribution $F_u(x)$ is


$$F_u(x) = \Pr(X - u \leq x | X > u)$$

- If u is VaR, then $F_u(x)$ is the probability that we exceed VaR by a particular amount (a shortfall) given that VaR is violated
- Key result is that as $u \rightarrow \infty$, $F_u(x)$ converges to the GPD, $G_{\xi, \beta}(x)$


- The GPD $G_{\xi,\beta}(x)$ is

$$G_{\xi,\beta}(x) = \begin{cases} 1 - \left(1 + \xi \frac{x}{\beta}\right)^{-\frac{1}{\xi}} & \xi \neq 0 \\ 1 - \exp\left(-\frac{x}{\beta}\right) & \xi = 0 \end{cases}$$

where $\beta > 0$ is the scale parameter; $x \geq 0$ when $\xi \geq 0$ and $0 \leq x \leq -\frac{\beta}{\xi}$ when $\xi < 0$

- We therefore need to estimate both shape(ξ) and scale(β) parameters when applying GDP
- Recall, for certain values of ξ the shape parameters, $G_{\xi,\beta}(\cdot)$ becomes one of the three distributions 

GEV and GPD

- The GEV is the limiting distribution of normalised maxima, whereas the GPD is the limiting distribution of normalised data beyond some high threshold 
- Note, the tail index is the same for both GPD and GEV distributions
- The parameters of GEV can be estimated from the log-likelihood function of GPD

VaR Under GPD

The VaR in the GPD case is:

$$\text{VaR}(p) = u + \frac{\beta}{\xi} \left[\left(\frac{1-p}{F(u)} \right)^{-\xi} - 1 \right]$$

Hill Method

- Alternatively, we could use the semi-parametric Hill estimator for the tail index in distribution $F(x) \approx 1 - Ax^{-\iota}$:

$$\hat{\xi} = \frac{1}{\hat{\iota}} = \frac{1}{C_T} \sum_{i=1}^{C_T} \log \frac{x_{(i)}}{u}$$

where $x_{(i)}$ is the notation of sorted data, for example, maxima is denoted as $x_{(1)}$

- As $T \rightarrow \infty$, $C_T \rightarrow \infty$ and $C_T/T \rightarrow 0$
- Note that the Hill estimator is sensitive to the choice of threshold, u

Which Method to Choose?

- *GPD*, as the name suggests, is more general and can be applied to all three types of tails
- *Hill method* on the other hand is in the maximum domain of attraction (MDA) of the Fréchet distribution
- Hence Hill method is only valid for fat-tailed data

Risk Analysis

- After estimation of the tail index, the next step is to apply a risk measure
- The problem is finding $\text{VaR}(p)$ such that

$$\Pr[X \leq -\text{VaR}(p)] = F_X(-\text{VaR}(p)) = p$$

where $F_X(u)$ is the probability of being in the tail, that is the returns exceeding the threshold u

Risk Analysis

- Let G be the distribution of X since we are in the left tail (ie $X \leq -u$). By the Pareto assumption we have:

$$G(-\text{VaR}(p)) = \left(\frac{\text{VaR}(p)}{u} \right)^{-\iota}$$

- And by the definition of conditional probability:

$$G(-\text{VaR}(p)) = \frac{p}{F_X(u)}$$

VaR Estimator

- Equating the previous two relationship, we obtain:

$$\text{VaR}(p) = u \left(\frac{F_X(u)}{p} \right)^{\frac{1}{\epsilon}}$$

- $F_X(u)$ can be estimated by the proportion of data beyond the threshold u , C_T/T
- The VaR estimator is therefore:

$$\widehat{\text{VaR}}(p) = u \left(\frac{C_T/T}{p} \right)^{\frac{1}{\epsilon}}$$

EVT Often Applied Inappropriately

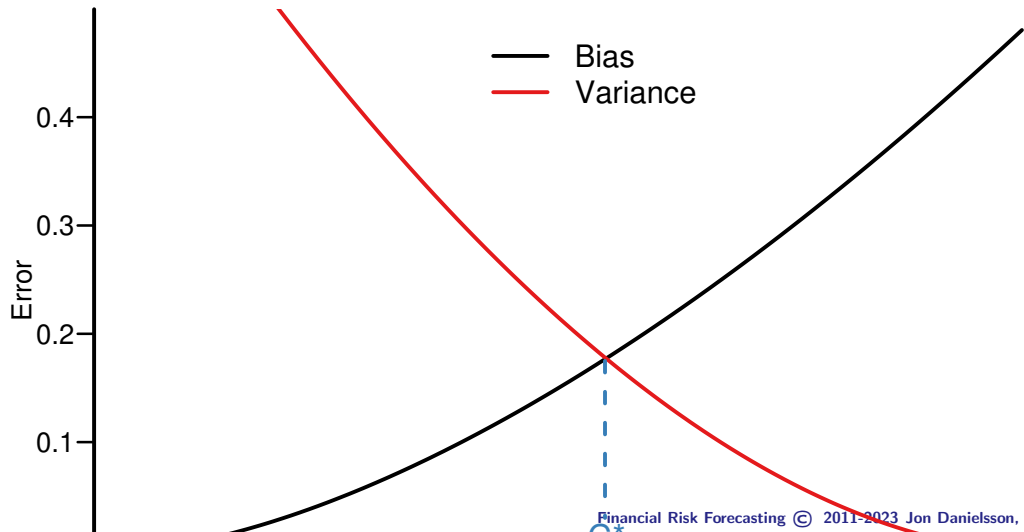
- EVT should only be applied in the tails
- The closer to the centre of the distribution, the more inaccurate the estimates are
- However, there are no rules to define when the estimates become inaccurate, it depends on the underlying distribution of the data
- In some cases, it may be accurate up to 1% or even 5%, while in other cases it is not reliable even up to 0.1%

Finding the Threshold

- Actual implementation of EVT is relatively simple and delivers good estimates where EVT holds
- The sample size T and the choice of probability level p depends on the underlying distribution of the data
- As a *rule of thumb*: $T \geq 1000$ and $p \leq 0.4\%$
- For applications with smaller sample sizes or less extreme probability levels, other techniques should be used
 - Such as HS or fat-tailed GARCH

- It can be challenging to estimate EVT parameters given the *effective sample size* is small
- This relates to choosing the number of observations in the tail, C_T
- We have 2 conflicting directions:
 1. By lowering C_T , we can reduce the estimation bias
 2. On the other hand, by increasing C_T , we can reduce the estimation variance

Optimal Threshold C_T^*

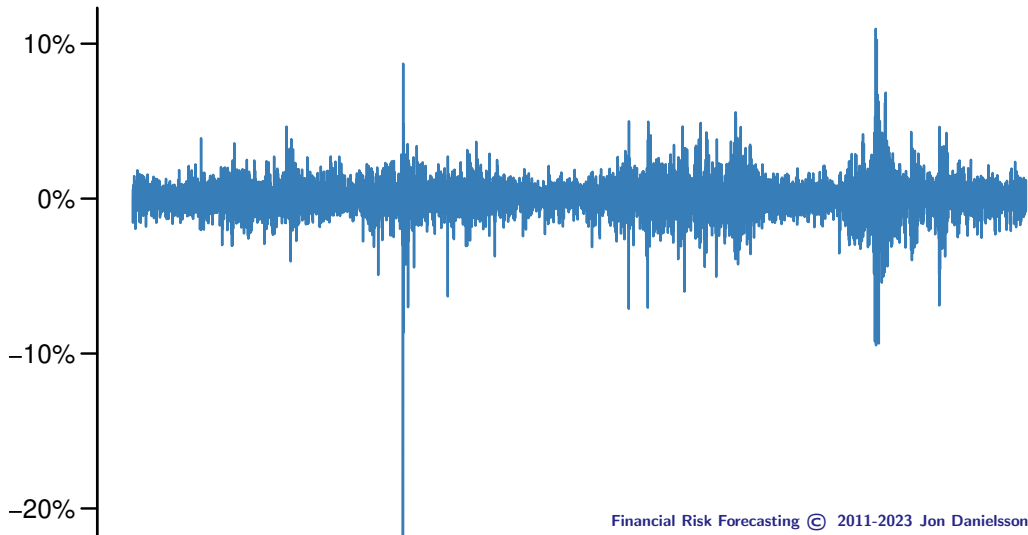


Optimal Threshold C_T^*

- If the underlying distribution is known, then deriving the optimal threshold is easy, but in such a case EVT is superfluous
- Most common approach to determine the optimal threshold is the *eyeball method* where we look for a region where the tail index seems to be stable
- More formal methods are based on minimising the mean squared error (MSE) of the Hill estimator, but such methods are not easy to implement

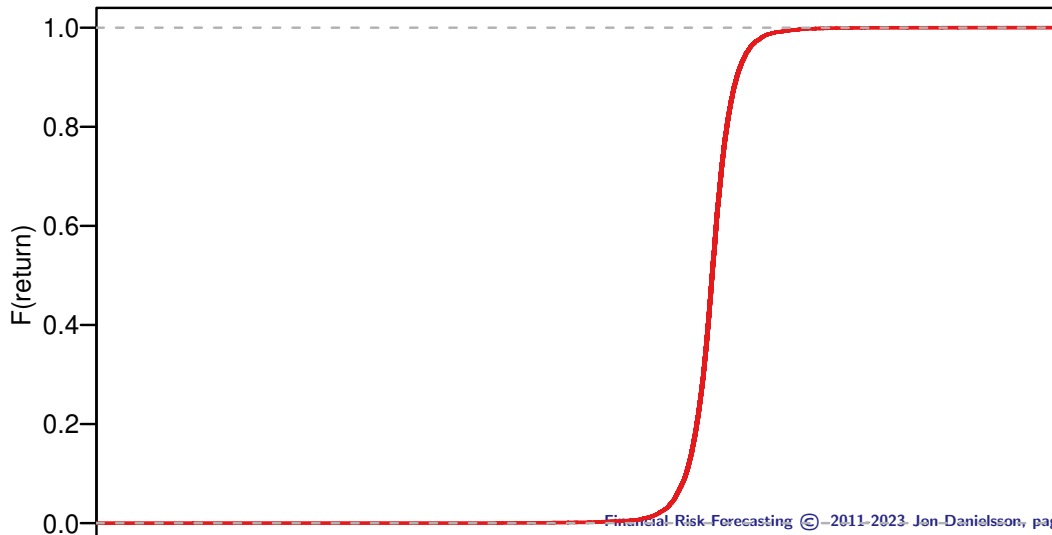
Application to the S&P-500 Index

Returns from 1975 to 2015 – 10,000 observations



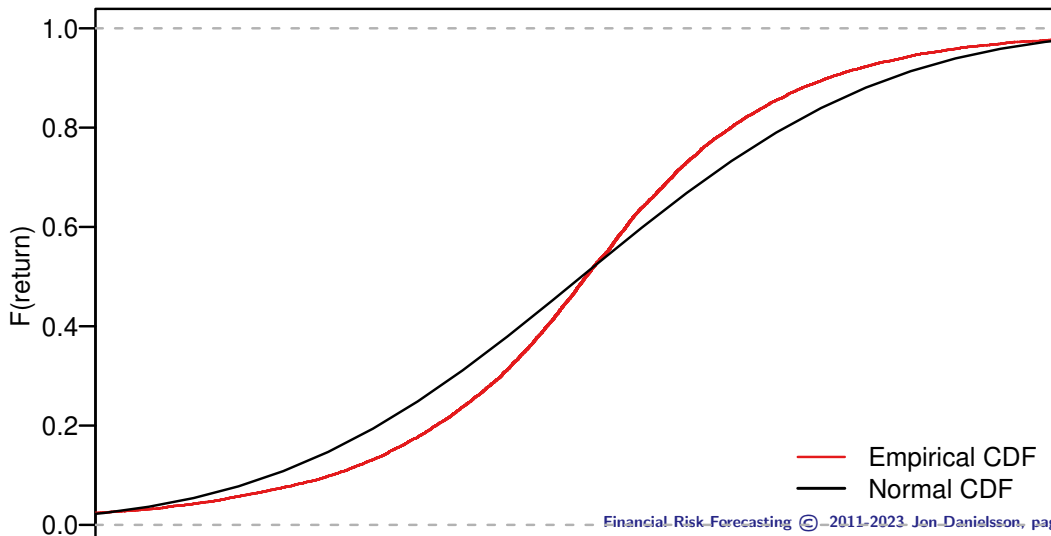
Distribution of S&P-500 Returns

Empirical distribution



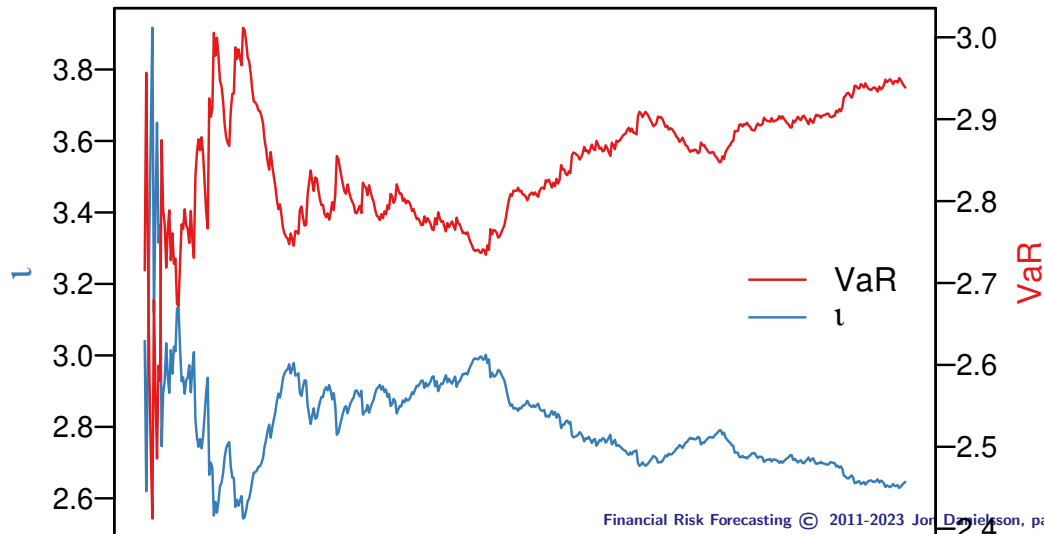
Distribution of S&P-500 Returns

Tails truncated



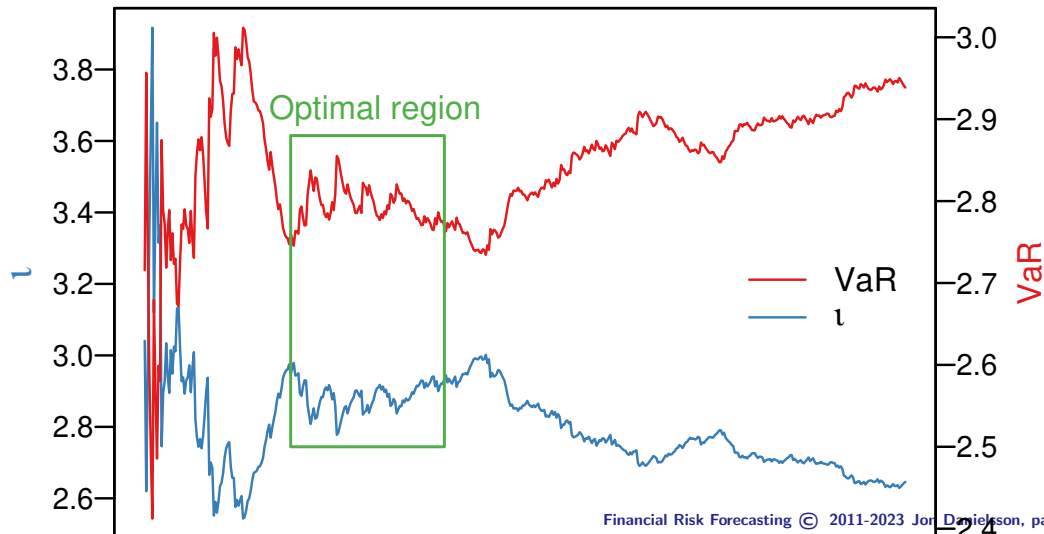
Hill Plot for Daily S&P-500 Returns

From 1975 to 2015



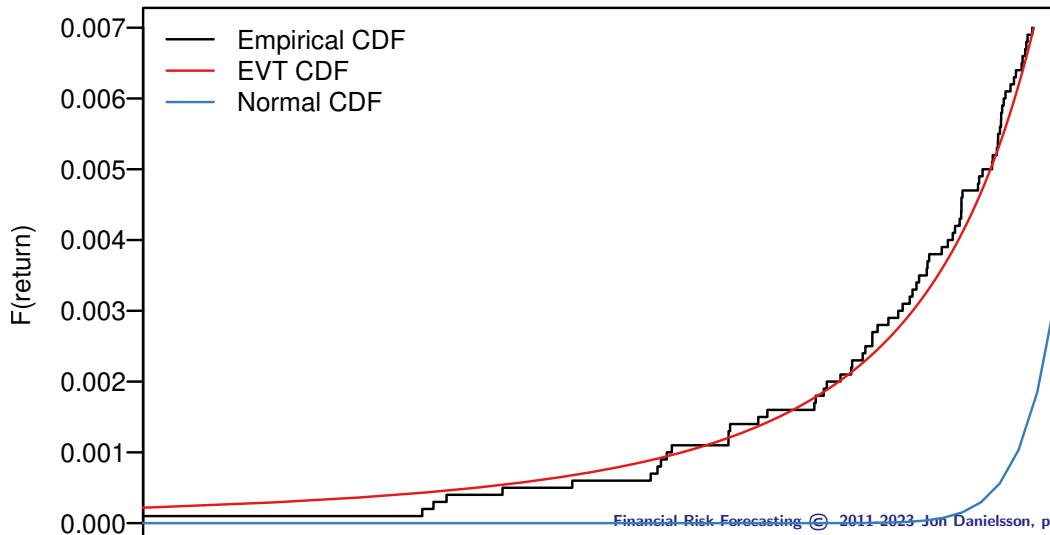
Hill Plot for Daily S&P-500 Returns

From 1975 to 2015



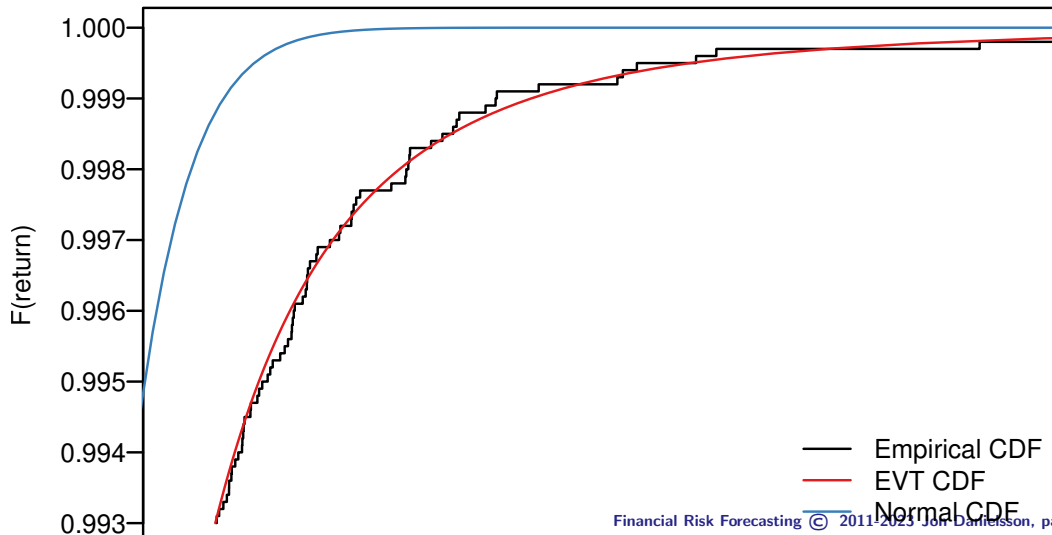
Upper and Lower Tails

The lower tail



Upper and Lower Tails

The upper tail



Aggregation and Convolution

Aggregation of Outcomes

- The act of adding up observations across time is known as *time aggregation*
- And the act of adding up observations across assets/portfolios is termed *convolution*

Feller 1971

Theorem Let X_1 and X_2 be two independent random variables with distribution functions satisfying

$$1 - F_i(x) = \mathbb{P}\{X_i > x\} \approx A_i x^{-\ell_i} \quad i = 1, 2$$

when $x \rightarrow \infty$. Note, A_i is a constant

Then, the distribution function F of the variable $X = X_1 + X_2$ in the positive tail can be approximated by 2 cases



Case 1 When $\iota_1 = \iota_2$ we say that the random variables are first-order similar and we set $\iota = \iota_1 = \iota_2$ and F satisfies

$$1 - F(x) = \mathbb{P}\{X > x\} \approx (A_1 + A_2)x^{-\iota}$$

Case 2 When $\iota_1 \neq \iota_2$ we set $\iota = \min(\iota_1, \iota_2)$ and F satisfies

$$1 - F(x) = \mathbb{P}\{X > x\} \approx Ax^{-\iota}$$

where A is the corresponding constant

- As a consequence, if two random variables are *identically distributed*, the distribution function of the sum (Case 1) will be given by

$$\mathbb{P}\{X_1 + X_2 > x\} \approx 2Ax^{-\iota}$$

- Hence the probability doubles when we combine two observations from different days
- But if one observations comes from a fatter tailed distribution than the other, then only the heavier tail matters (Case 2)

Time Scaling

Theorem (de Vries 1998) Suppose X has finite variance with a tail index $\iota > 2$. At a constant risk level p , increasing the investment horizon from 1 to T periods increases the VaR by a factor:

$$T^{1/\iota}$$

Note, EVT distributions retain the same tail index for longer period returns

- Recall from chapter 4, under Basel Accords, financial institutions are required to calculate VaR for a 10-day holding periods
- The rules allow the 10-day VaR to be calculated by scaling the one-day VaR by $\sqrt{10}$
- The theorem shows that the scaling parameter is slower than the square-root-of-time adjustment
- Intuitively, as extreme values are more rare, they should aggregate at a slower rate than the normal distribution
- For example, if $\iota = 4$, $10^{1/\iota} = 1.78$, which is less than $\sqrt{10} = 3.16$

VaR and the Time Aggregation of Fat Tail Distributions

Risk level	5%	1%	0.5%	0.1%	0.05%	0.005%
<i>Extreme value</i>						
1 Day	0.9	1.5	1.7	2.5	3.0	5.1
10 Day	1.6	2.5	3.0	4.3	5.1	8.9
<i>Normal</i>						
1 Day	1.0	1.4	1.6	1.9	2.0	2.3
10 Day	3.2	4.5	4.9	5.9	6.3	7.5

- For one-day horizons, we see that in general EVT VaR is higher than VaR under normality, especially for more extreme risk levels
- This is balanced by the fact that 10-day EVT VaR is less than the normal VaR
- This seems to suggest that the square-root-of-time rule may be sufficiently prudent for longer horizons
- It is important to keep in mind that ι root rule (de Vries) only holds *asymptotically*

Time Dependence


Time Dependence

- Recall the assumption of IID returns in the section on EVT, which suggests that EVT may not be relevant for financial data
- Fortunately, we **do not need** an IID assumption, since EVT estimators are consistent and unbiased even in the presence of higher moment dependence
- We can explicitly model extreme dependence using the *extremal index*

Example

- Let us consider extreme dependence in a MA(1) process:

$$Y_t = X_t + \alpha X_{t-1} \quad |\alpha| < 1$$


- Let X_t and X_{t-1} be IID such that $\Pr\{X_t > x\} \rightarrow Ax^{-\ell}$ as $x \rightarrow \infty$. Then by Feller's theorem 

$$\mathbb{P}\{Y_t \geq x\} \approx (1 + \alpha^\ell)Ax^{-\ell} \quad \text{as } x \rightarrow \infty$$

- Dependence enters *“linearly”* by means of the coefficient α^ℓ . But the tail shape is unchanged
- This example suggest that time dependence has same effect as having an IID sample with fewer observations

- Suppose we record each observation twice:

$$Y_1 = X_1, Y_2 = X_1, Y_3 = X_2, \dots$$

- And it increases the sample size to $D = 2T$. Let us define $M_D \equiv \max(Y_1, \dots, Y_D)$. Evidently from Fisher-Tippett and Gnedenko theorem: 

$$\mathbb{P}\{M_D \leq x\} = F^T(x) = F^{\frac{D}{2}}(x)$$

supposing $a_T = 0$ and $b_T = 1$

- The important result here is that *dependence increases the probability that the maximum is below threshold x*

Extremal Index

Extremal index ψ It is a measure of tail dependence and $0 < \psi \leq 1$

- If the data are *independent* then we get

$$\mathbb{P}\{M_T \leq x\} \rightarrow e^{-x^{-\iota}} \quad \text{as } T \rightarrow \infty$$

when $a_T = 0$ and $b_T = 1$

- If the data are *dependent*, the limit distribution is

$$\mathbb{P}\{M_D \leq x\} \rightarrow \left(e^{-x^{-\iota}}\right)^{\psi} = e^{-\psi x^{-\iota}}$$

- $\frac{1}{\psi}$ is a measure of the *cluster size* in large samples, for double-recorded data $\psi = \frac{1}{2}$
- For the MA(1) process in the previous example, we obtain the following

$$\mathbb{P} \left\{ T^{-\frac{1}{\psi}} M_D \leq x \right\} \rightarrow \exp \left(-\frac{1}{1 + \alpha^\psi} x^{-\psi} \right)$$

where $\psi = \frac{1}{1 + \alpha^\psi}$

Dependence in ARCH

- Consider the normal ARCH(1) process:

$$Y_t = \sigma_t Z_t$$

$$\sigma_t^2 = \omega + \alpha Y_{t-1}^2$$

$$Z_t \sim \mathcal{N}(0, 1)$$

- Subsequent returns are uncorrelated but are *not independent*, since

$$\text{Cov}(Y_t, Y_{t-1}) = 0$$

$$\text{Cov}(Y_t^2, Y_{t-1}^2) \neq 0$$

- Even when Y_t is conditionally normally distributed, we noted in chapter 2 that the unconditional distribution of Y is fat tailed
- de Haan et al. show that the unconditional distribution of Y is given by

$$\Gamma\left(\frac{\iota}{2} + \frac{1}{2}\right) = \sqrt{\pi}(2\alpha)^{-\iota/2}$$

Extremal Index for ARCH(1) – Example

- Extremal index for the ARCH(1) process can be solved using the previous equation
- From the table below, we see that the higher the α , the fatter the tails and the higher the level of clustering

α	0.10	0.50	0.90	0.99
ι	26.48	4.73	2.30	2.02
ψ	0.99	0.72	0.46	0.42

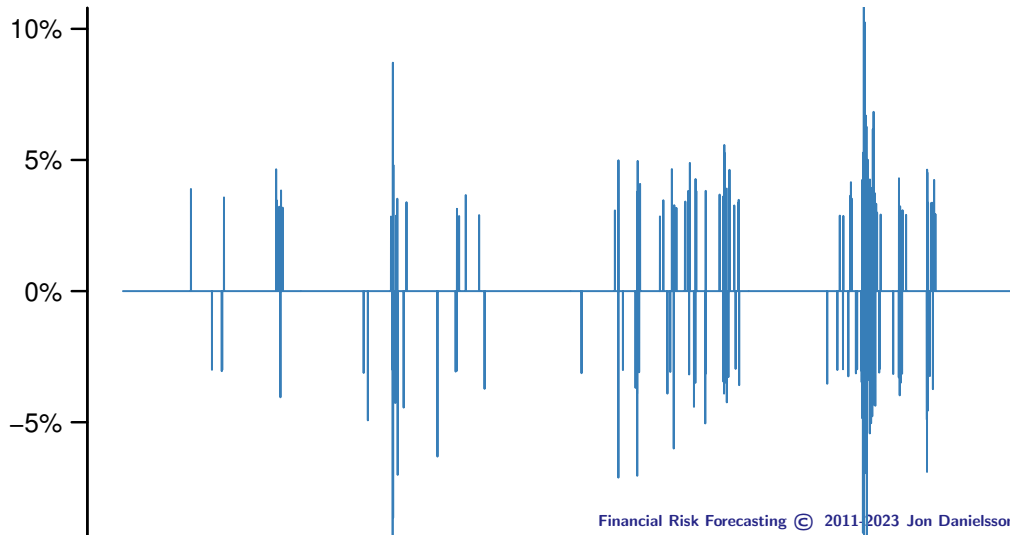
- Similar results can be obtained for GARCH

When Does Dependence Matter?

- The importance of extreme dependence and the extremal index ψ depends on the underlying applications
- Dependence can be *ignored* if we are dealing with *unconditional probabilities*
- And dependence *matters* when calculating *conditional probabilities*
- For many stochastic processes, including GARCH, the time between tail events become increasingly independent

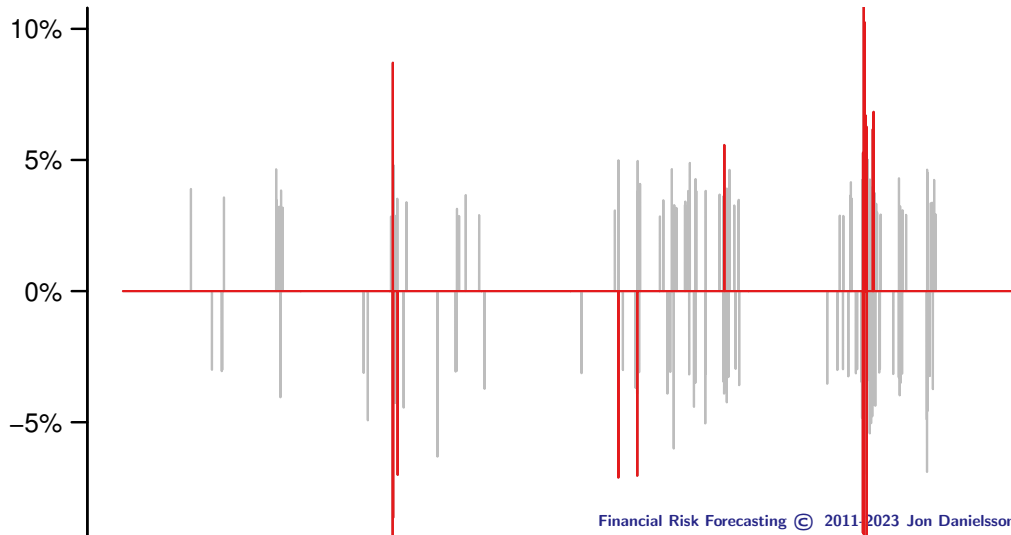
Example – S&P-500 Index Extremes

From 1970 to 2015, 1% events



Example – S&P-500 Index Extremes

From 1970 to 2015, 0.1% events



Example – S&P-500 Index Extremes

0.1% events during the crisis

